

Robust PCA for High-Dimensional Data

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Talk by Shie Mannor, The Technion
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Thank you for staying for the graveyard session

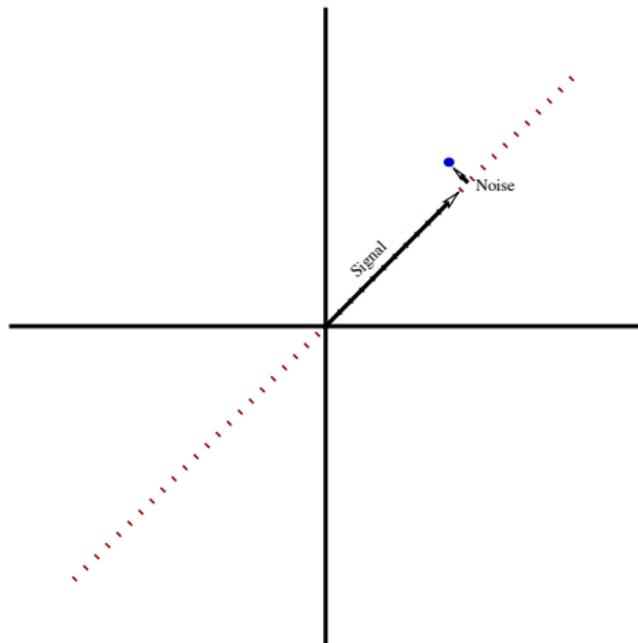
PCA - in Words

- Observe high-dimensional points
- Find least-square-error subspace approximation

- Many applications in feature-extraction and compression
 - data analysis
 - communication theory
 - pattern recognition
 - image processing

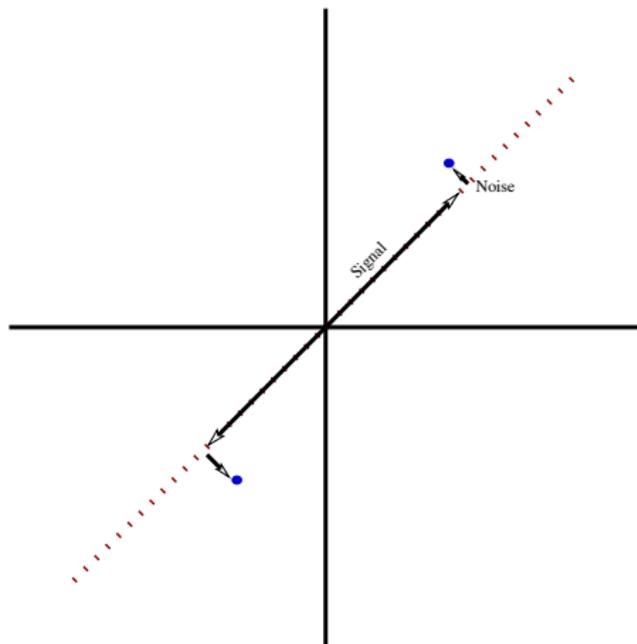
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Observe points: $\mathbf{y} = \mathbf{A}\mathbf{x} + \mathbf{v}$.



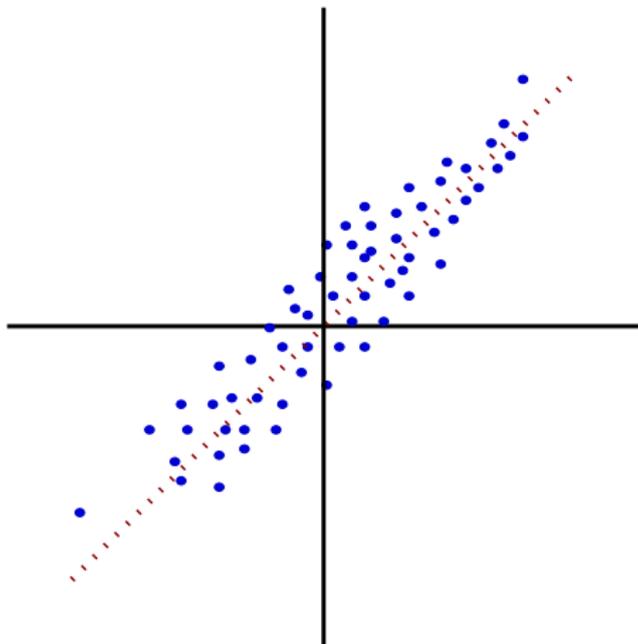
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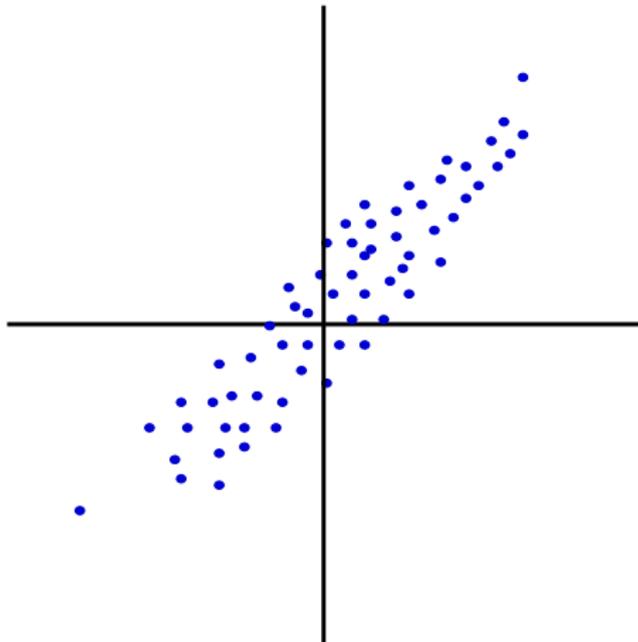
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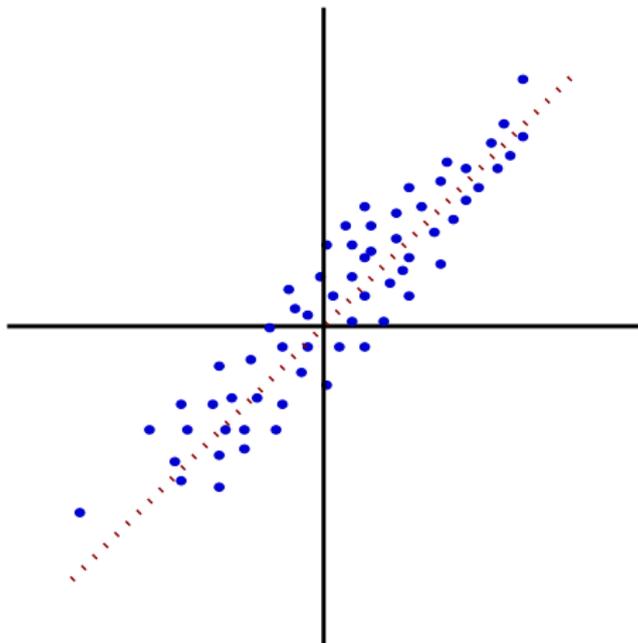
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Goal: Find least-square-error subspace approximation.



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PCA - in Math

- Least-square-error subspace approximation
- How: Singular value decomposition (SVD) performs eigenvector decomposition of the sample-covariance matrix
- Magic of SVD: solving a non-convex problem
- Cannot replace quadratic objective here.
- Consequence: Sensitive to outliers
 - Even **one** outlier can make the output arbitrarily skewed;
 - What about a constant fraction of “outliers”?

This Talk: High Dimensions and Corruption

Two key differences to pictures shown

- (A) High-dimensional regime: # observations \leq dimensionality.
- (B) A constant fraction of points arbitrarily corrupted.

Outline

1. **Motivation: PCA, High dimensions, corruption**
2. Where things get tricky: usual tools fail
3. HR-PCA: the algorithm
4. The Proof Ideas (and some details)
5. Conclusion

High-Dimensional Data

- What is high-dimensional data:
 $\# \text{dimensionality} \approx \# \text{ observations}$.
- Why high-dimensional data analysis:
 - Many practical examples: DNA microarray, financial data, semantic indexing, images, etc

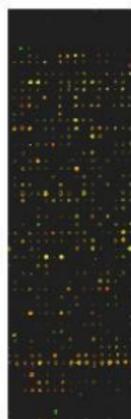


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MicroArray:
24,401 dim.

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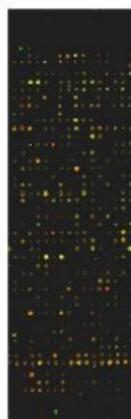


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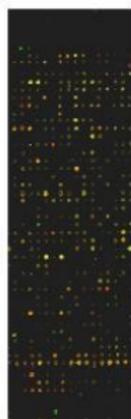


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 - The **kernel trick** generates high-dimensional data
 - Traditional statistical tools do not work

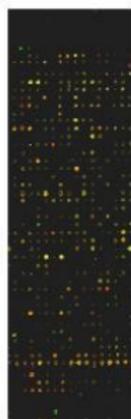


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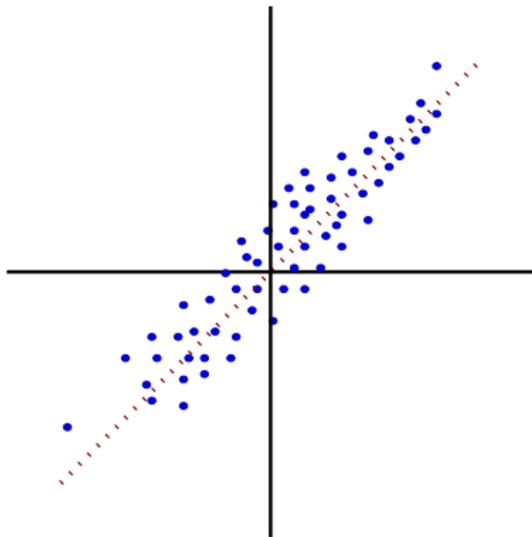


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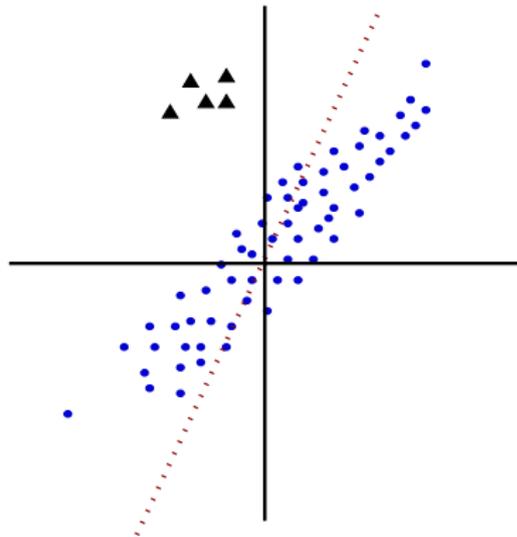


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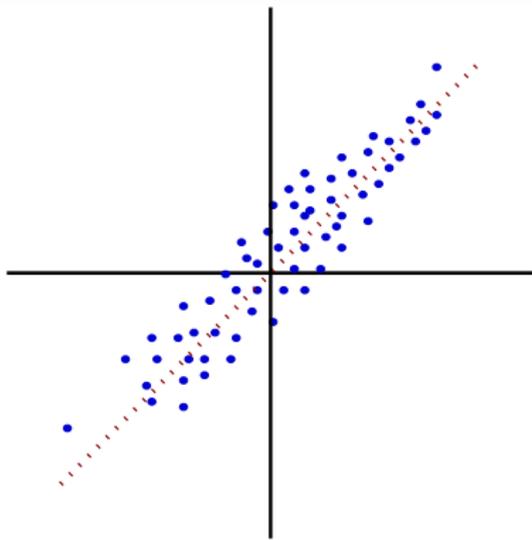


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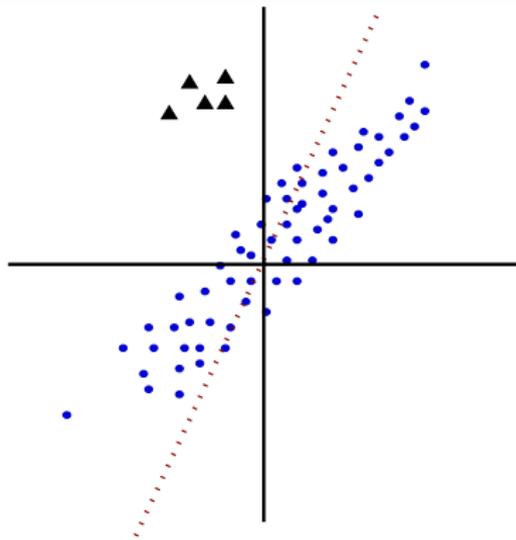


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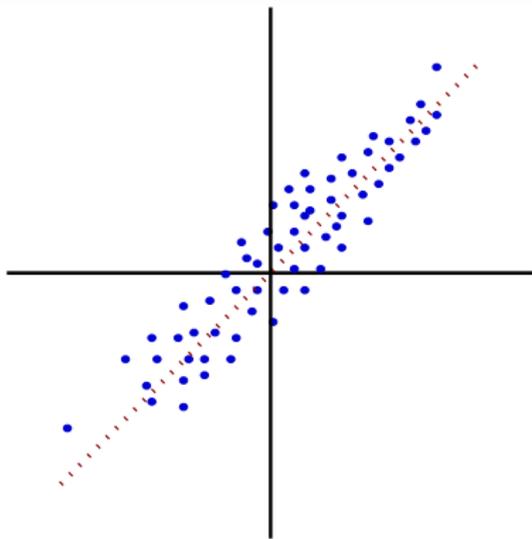


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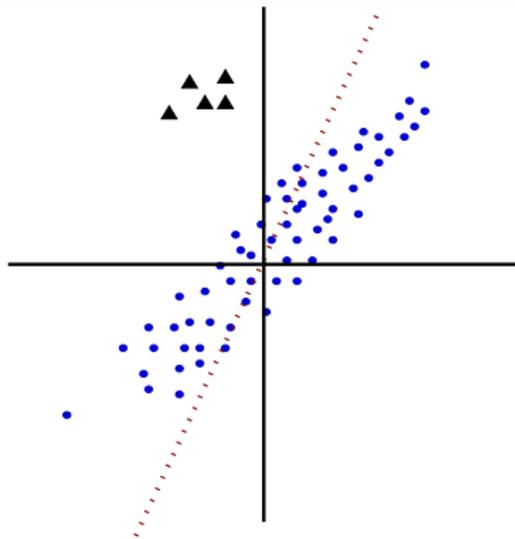


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Our Goal: Robust PCA

- Want robustness to arbitrarily corrupted data.
- One measure: Breakdown point
- Instead: bounded error measure between true PCs and output PCs.
- Bound will depend on:
 - Fraction of outliers.
 - Tails of true distribution.

Problem Setup

- “Authentic Samples” $\mathbf{z}_1, \dots, \mathbf{z}_t \in \mathbb{R}^m$: $\mathbf{z}_j = \mathbf{A}\mathbf{x}_j + \mathbf{n}_j$,

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- Regime of interest:
 - $n \approx m \gg d$
 - $\sigma = \|\mathbf{A}^\top \mathbf{A}\| \gg 1$ (scales slowly).
- Objective: Retrieve \mathbf{A}

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Features of the High Dimensional regime

- Noise Explosion in High Dimensions: noise magnitude scales faster than the signal noise;
- SNR goes to zero
 - If $\mathbf{n} \sim N(0, I_m)$, then $\mathbb{E}\|\mathbf{n}\|_2 = \sqrt{m}$, with very sharp concentration.
 - Meanwhile: $\mathbf{E}\|A\mathbf{x}\|_2 \leq \sigma\sqrt{d}$.
- Consequences:
 - Magnitude of true samples may be much bigger than outlier magnitude.
 - The direction of each sample will be approximately orthogonal to the direction of the signal;

Features of the High Dimensional regime: Pictures

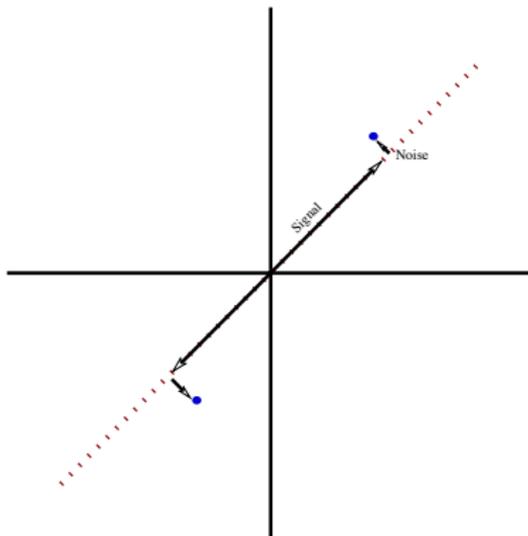


Figure: Recall low-dimensional regime

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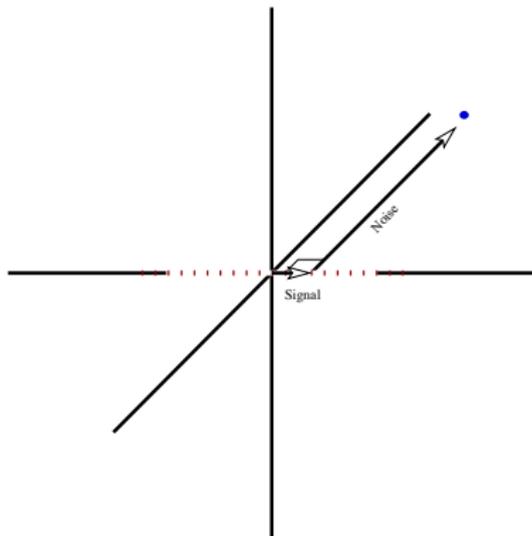


Figure: High dimensions are different: Noise \gg Signal

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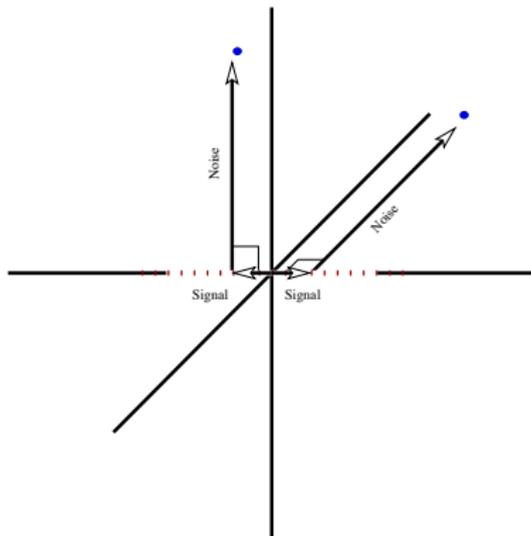


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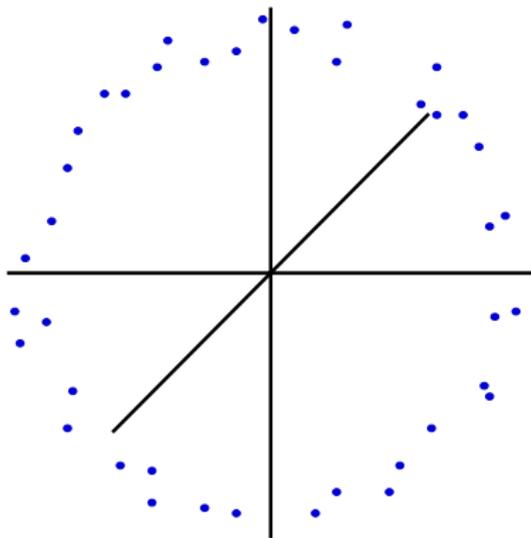


Figure: Every point equidistant from origin and from other points!

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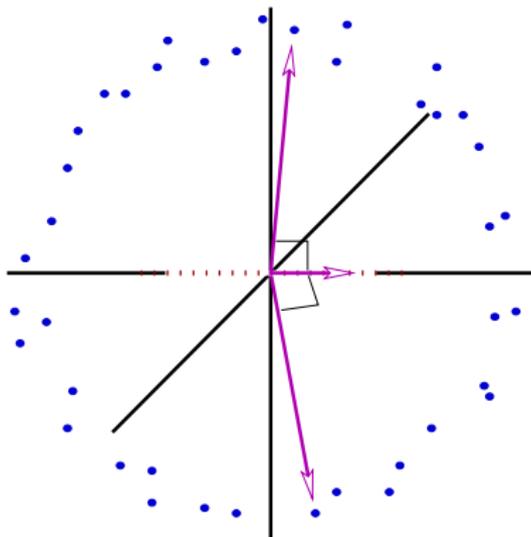


Figure: And every point perpendicular to signal space

Trouble in High Dimensions

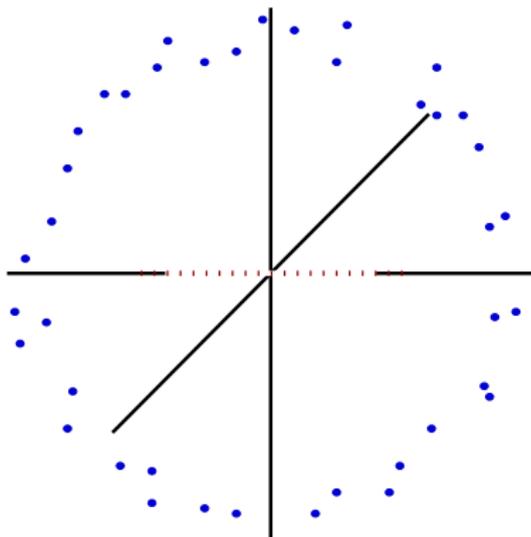
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- Some approaches that will not work:
- Leave-one-out (more generally, subsample, compare):
 - Either sample size very small: problem
 - or
 - Have many corrupted points in each subsample: problem
- Standard Robust PCA: PCA on a robust estimation of the covariance
 - Consistency requires $\#(\text{observations}) \gg \#(\text{dimension})$
 - Not enough observations in high-dimensional case

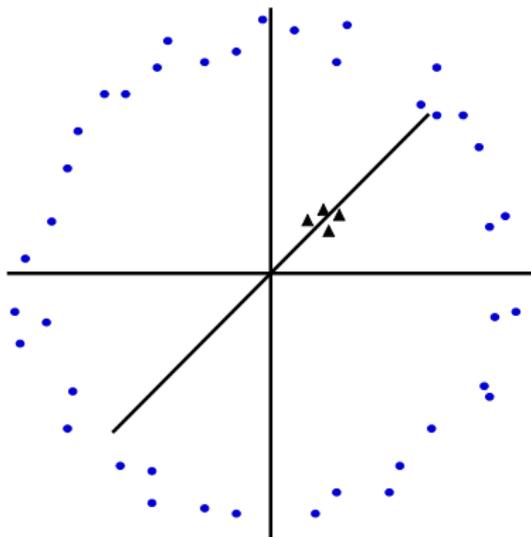
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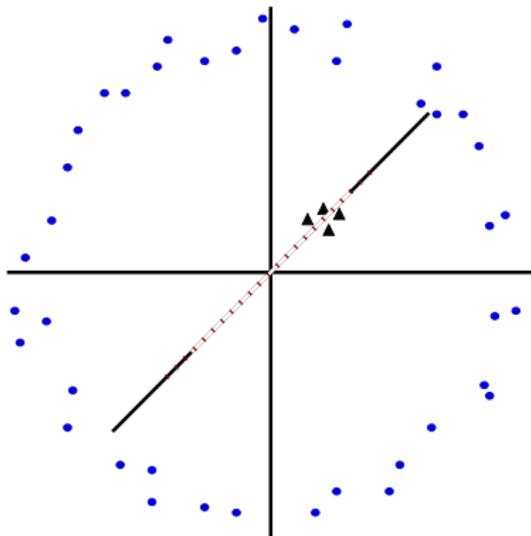
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- Remove points with large Mahalanobis distance
 - Same example: All λn corrupted points: aligned, length $O(\sigma) \ll \sqrt{m}$.
 - Very large impact on PCA output.
 - But: Mahalanobis distance of outliers very small.

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 - Same example: All λn corrupted points: aligned, length $O(\sigma) \ll \sqrt{m}$.
 - Very large impact on PCA output.
 - But: Mahalanobis distance of outliers very small.
- Remove points with large Stahel-Donoho distance

$$u_i \triangleq \sup_{\|\mathbf{w}\|=1} \frac{|\mathbf{w}^\top \mathbf{y}_i - \text{med}_j(\mathbf{w}^\top \mathbf{y}_j)|}{\text{med}_k |\mathbf{w}^\top \mathbf{y}_k - \text{med}_j(\mathbf{w}^\top \mathbf{y}_j)|}.$$

- Same example: impact large, but Stahel-Donoho outlyingness small.

Trouble in High Dimensions

- For these reasons: Some robust covariance estimators have breakdown point = $O(1/m)$, m = dimensions.
 - M-estimator,
 - Convex peeling, Ellipsoidal Peeling,
 - Classical outlier rejection
 - Iterative deletion, iterative trimming,
 - and others...
- These approaches cannot work in high-dimensional regime.

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- Minimum volume ellipsoid; Minimum covariance determinant:
 - Ill-posed: many zero-volume ellipsoids containing data
 - Intractable: removing a fraction of points combinatorial.
- Projection pursuit – maximize univariate estimator
 - Problems are non-convex: Intractable.
 - Choosing subset of directions generated by points: authentic points \perp to signal space, hence no good in high dimensions.

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High-dimensional Robust PCA: Main Idea

- Get candidate directions from standard PCA (get \mathbf{w}).
- Project, and use a robust variance estimator: variance of points nearer origin.
 - Outliers can be near origin. But: impact controlled.
- Random removal of “strange” points.

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- Desired properties of an algorithm:
 - Tractable (same complexity as standard PCA);
 - Robust to outliers: performance guarantees;
 - Asymptotically optimal: $t = o(n)$ perfect recovery.
 - Easily kernelizable;

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- Assumptions:
 - n, m scale to infinity together;
 - $\sigma = \|\mathbf{A}^\top \mathbf{A}\|$ “big” (scales to infinity slowly);
 - μ : spherically symmetric; abs continuous; exponential tails.

Objective & Performance Measurement

- For output PCs $\mathbf{w}_1, \dots, \mathbf{w}_d$, “Expressed Variance” w.r.t. $\mathbf{w}_1^{\text{true}}, \dots, \mathbf{w}_d^{\text{true}}$

$$E_V(\mathbf{w}_1, \dots, \mathbf{w}_d) \triangleq \frac{\sum_{i=1}^d \mathbf{w}_i^{\top} \mathbf{A} \mathbf{A}^{\top} \mathbf{w}_i}{\sum_{i=1}^d (\mathbf{w}_i^{\text{true}})^{\top} \mathbf{A} \mathbf{A}^{\top} \mathbf{w}_i^{\text{true}}} \leq 1.$$

- $E_V = 1$ if the subspace spanned by true PCs is recovered.
- For $d = 1$, $E_V(\mathbf{w}_1) = \cos^2(\angle \mathbf{w}_1, \mathbf{w}_1^{\text{true}})$.

A Robust Variance Estimator

- **Robust Variance Estimator:** $\bar{V}_{\hat{t}}(\mathbf{w}) \triangleq \frac{1}{n} \sum_{i=1}^{\hat{t}} |\mathbf{w}^\top \mathbf{y}|_{(i)}^2$.
- **Order statistics:** $\alpha_1, \dots, \alpha_n \in \mathbb{R}$, then $\alpha_{(1)} \leq \alpha_{(2)} \leq \dots \leq \alpha_{(n)}$.
- **Idea:** If outliers small, their impact is controlled.

The HR-PCA Algorithm

- (1) Perform PCA on empirical covariance.
- (2) If robust variance estimate in PC directions highest yet, record it, and PCs.
- (3) Randomly remove a point in proportion to its variance along PCs.
- (4) Repeat until "enough" points removed.
- (5) Output the last PCs recorded.

The HR-PCA Algorithm

- (1) Perform PCA on empirical covariance: $\{\mathbf{w}_1, \dots, \mathbf{w}_d\}$.
- (2) Compute $b = \text{RVE}(\{\mathbf{w}_1, \dots, \mathbf{w}_d\})$. If $b > b^*$,
 - Update $b^* = b$
 - Update $\{\mathbf{w}_1^*, \dots, \mathbf{w}_d^*\} = \{\mathbf{w}_1, \dots, \mathbf{w}_d\}$.
- (3) Randomly remove a point in proportion to its variance along PCs.
- (4) Repeat until all points removed.
- (5) Output the last PCs recorded: $\{\mathbf{w}_1^*, \dots, \mathbf{w}_d^*\}$.

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- Things that can go wrong:
 - * Remove authentic points
 - * May not ultimately report “best outcome.”
 - * Corrupted points may contribute to ultimately reported PCs.
- But: we show the error due to all such factors is controlled.

The Guarantees: Finite Sample + Asymptotic

- Results will depend on:
 - Fraction of outliers: λ .
 - Tails of μ .
- Define: $\mathcal{V} : [0, 1] \rightarrow [0, 1]$

$$\mathcal{V}(\alpha) = \int_{-c_\alpha}^{c_\alpha} x^2 \bar{\mu}(dx).$$

The Guarantees: Finite Sample + Asymptotic

Theorem: The following holds in probability (n, m, σ scale):

$$\text{E.V.}(\text{output}) \geq \max_{\kappa} \left[\frac{\mathcal{V} \left(1 - \frac{\lambda^*(1+\kappa)}{(1-\lambda^*)\kappa} \right)}{(1+\kappa)} \right] \times \left[\frac{\mathcal{V} \left(\frac{\hat{t}}{t} - \frac{\lambda^*}{1-\lambda^*} \right)}{\mathcal{V} \left(\frac{\hat{t}}{t} \right)} \right].$$

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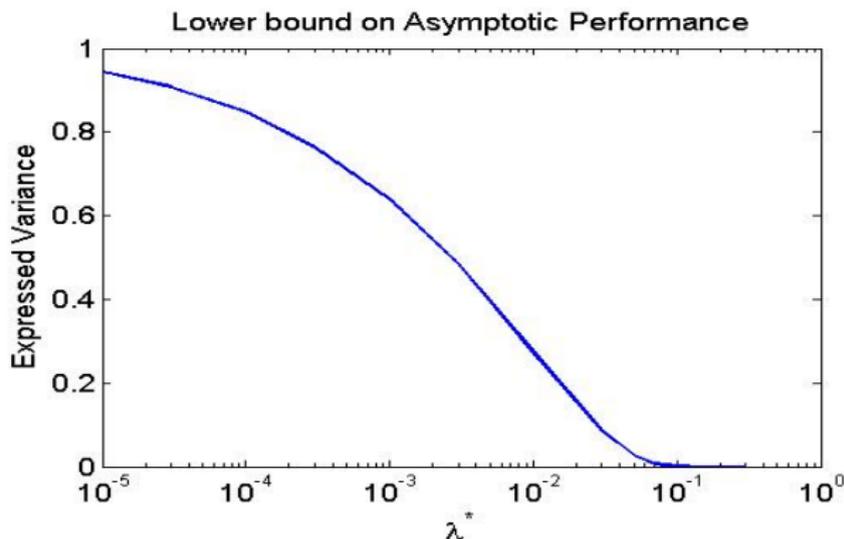
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- The Bound:
 - Term 1: May not remove all outliers, and some authentic points may be removed.
 - Term 2: May have small outliers that alter PC directions.
- If $t = o(n)$, RHS = 1: optimal recovery.
- Breakdown point: 1/2.

Asymptotic Performance Guarantee

E.V. is lower bounded by



If the *proportion* of outliers goes to zero: the Expressed Variance equals 1.

Proof Idea

- (1) “Blessing of dimensionality”: empirical covariance estimates good, even for high-dimensional regime;
- (2) Random removal: have a “good” solution, or outlier is removed with large probability;
- (3) Therefore: at some early iteration, algorithm finds a “good” solution.
- (4) Output of algorithm has higher robust variance estimate than the “good” solution. We show output must then also be (almost as) “good.”

Proof Idea - Step 1

With high probability:

- (1.a) Largest eigenvalue of the empirical noise covariance matrix is bounded:

$$\sup_{\mathbf{w} \in \mathcal{S}_m} \frac{1}{n} \sum_{i=1}^t (\mathbf{w}^\top \mathbf{n}_i)^2 \leq c.$$

- (1.b) Largest eigenvalue of the signals in original space converges to 1:

$$\sup_{\mathbf{w} \in \mathcal{S}_d} \left| \frac{1}{t} \sum_{i=1}^t (\mathbf{w}^\top \mathbf{x}_i)^2 - 1 \right| \leq \epsilon.$$

Proof Idea - Step 1

(1.c) RVE is a valid variance estimator for the d -dimensional signals \mathbf{x} :

$$\sup_{\mathbf{w} \in \mathcal{S}_d} \left| \frac{1}{\hat{t}} \sum_{i=1}^{\hat{t}} |\mathbf{w}^\top \mathbf{x}|_{(i)}^2 - \mathcal{V} \left(\frac{\hat{t}}{\hat{t}} \right) \right| \leq \epsilon.$$

(1.d) RVE is a valid estimator of the variance of the authentic samples, $\mathbf{z} = \mathbf{A}\mathbf{x} + \mathbf{n}$: uniformly over all $\mathbf{w} \in \mathcal{S}_m$,

$$(1 - \epsilon) \|\mathbf{w}^\top \mathbf{A}\|^2 \mathcal{V} \left(\frac{t'}{t} \right) - c \|\mathbf{w}^\top \mathbf{A}\| \leq \frac{1}{t} \sum_{i=1}^{t'} |\mathbf{w}^\top \mathbf{z}|_{(i)}^2 \leq$$

$$(1 + \epsilon) \|\mathbf{w}^\top \mathbf{A}\|^2 \mathcal{V} \left(\frac{t'}{t} \right) + c \|\mathbf{w}^\top \mathbf{A}\|.$$

Proof - Step 1.a - details

(1.a) Largest eigenvalue of the variance of noise matrix is bounded:

$$\sup_{\mathbf{w} \in \mathcal{S}_m} \frac{1}{n} \sum_{i=1}^t (\mathbf{w}^\top \mathbf{n}_i)^2 \leq c.$$

- Two keys: “blessing of dimensionality” and uniform laws of large numbers.

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- Two keys: “blessing of dimensionality” and uniform laws of large numbers.
- Step 1 (a): Need basic Lemma:
- **Lemma:** For Γ a $m \times t$ matrix ($m \leq t$), $\Gamma_{ij} \sim \mathcal{N}(0, 1)$, i.i.d.:

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Proof - Step 1.a - details

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- **Observation:**

$$\sup_{\mathbf{w} \in \mathcal{S}_m} \frac{1}{t} \sum_{i=1}^t (\mathbf{w}^\top \mathbf{n}_i)^2 = \lambda_{\max}(\Gamma\Gamma^\top)/t = \sigma_{\max}^2(\Gamma)/t.$$

Proof - Step 1.a - An Aside

- Where do these results come from:
- Basic idea: *dimension-free* concentration of measure
- **Theorem:** Let F be L -Lipschitz w.r.t. Euclidean norm, $X \sim N(0, I)$ standard Gaussian measure. M_F the mean of $F(X)$. Then

$$\mathbb{P}(F(X) \geq M_F + \xi) \leq e^{-\xi^2/2L^2}.$$

- Basic observation: $\sigma_{\max}(\cdot) : \mathbb{R}^{n_1 \times n_2} \rightarrow \mathbb{R}$ is 1-Lipschitz.
- Two nice references: (a) Davidson and Szarek: Operators, Random Matrices & Banach Spaces; (b) Matousek: Lectures on Discrete Geometry.

Proof Idea

- (1) “Blessing of dimensionality”: empirical covariance estimates good, even for high-dimensional regime;
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- Hence: $\{\mathbf{w}_1, \dots, \mathbf{w}_d\}$ must be close to true PCs.
- **Theorem:** If $\mathcal{G}^c(s)$ — step s is not good — then next point removed is an outlier with probability at least $\frac{\kappa}{1+\kappa}$.

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- Intuition: Suppose subsequent steps were independent.
 - Since, “expected number of corrupted points removed each step” is $\kappa/(1 + \kappa)$.
 - After M steps, expected corrupted points removed is $M \frac{\kappa}{1+\kappa}$.
 - Therefore: All the outliers removed after $M = \lambda n \frac{1+\kappa}{\kappa} (1 + \varepsilon)$ steps, with exponentially high probability.

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 - The Problem: not i.i.d.
 - The Fix: use martingales and Azuma-Hoeffding.

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- Azuma-Hoeffding completes the proof.

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Proof Idea - Step 4

- Putting it all together:
- An early iteration produces directions $\hat{\mathbf{w}}_1, \dots, \hat{\mathbf{w}}_d$ that have “most of” the variance.
- Bound quality on these directions:

$$E_V(\hat{\mathbf{w}}_1, \dots, \hat{\mathbf{w}}_d) \triangleq \frac{\sum_{i=1}^d \hat{\mathbf{w}}_i^\top \mathbf{A} \mathbf{A}^\top \hat{\mathbf{w}}_i}{\sum_{i=1}^d (\mathbf{w}_i^{\text{true}})^\top \mathbf{A} \mathbf{A}^\top \mathbf{w}_i^{\text{true}}}.$$

- The final algorithm only produces directions $\mathbf{w}_1^*, \dots, \mathbf{w}_d^*$ with biggest robust variance estimator.
- Bound quality on these directions:

$$E_V(\mathbf{w}_1^*, \dots, \mathbf{w}_d^*) \triangleq \frac{\sum_{i=1}^d (\mathbf{w}_i^*)^\top \mathbf{A} \mathbf{A}^\top \mathbf{w}_i^*}{\sum_{i=1}^d \sum_{i=1}^d \hat{\mathbf{w}}_i^\top \mathbf{A} \mathbf{A}^\top \hat{\mathbf{w}}_i}.$$

Kernelization

- Using a kernel function $k(\cdot, \cdot)$ to represent a feature mapping $\Upsilon(\cdot)$
- PCA can be kernelized using Kernel PCA, with output in a form $\mathbf{v}_q = \sum_{i=1}^{n-s} \alpha_i(q) \Upsilon(\hat{\mathbf{y}}_i)$, $q = 1, \dots, d$.
- HR-PCA Algorithm requires:
 - Computing PCA;
 - Computing Robust Variance Estimator;
- Both steps can be done.

Conclusion

- Methodology for handling dimensionality reduction when:
 1. $\#(\text{Observation}) \sim \#(\text{Dimension})$
 2. $\#(\text{Outliers})$ is "large"
- The key idea: verify projections statistics behave in a certain way, if not - probabilistic point removal
- Works well in simulations

On the todo list:

- Generalize to other identification problems with outliers: when a probabilistic model is available
- Extend to stochastic programming with corrupted sampled data
- Looking for an online algorithm.